

Meta-analysis demonstrates association of the TGF-β1 gene -C509T polymorphism with susceptibility to IgA nephropathy in European but not in Asian populations

H. Wang^{1,2}, P. Li² and Z.-C. Feng¹

¹School of Medicine and Health Management, HuaZhong University of Science and Technology, Wuhan, China ²First Affiliated Hospital, Anhui Medical University, Hefei, Anhui, China

Corresponding author: Z.-C. Feng E-mail: zhanchunfengcn@163.com

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ABSTRACT. There are conflicting reports associating the TGF-B1 gene -C509T polymorphism with susceptibility to IgA nephropathy. We investigated this association through a meta-analysis. Case-control studies were searched up to January 2012; the genotype frequencies in the control group were found to be consistent with Hardy-Weinberg equilibrium. Publication bias was tested by funnel plot and the Egger regression test. Eight studies, comprising 1364 cases and 1483 controls, were included. Significant heterogeneity was observed ($\chi^2 = 18.29$, P = 0.01). Under the random-effects model, the overall odds ratio (OR) was 1.01 [95% confidence interval (95%CI) = 0.74-1.38; P = 0.94]. In the subgroup analysis based on ethnicities, no significant effect was observed in the Asian descent groups (five comparisons, OR = 0.78; 95%CI=0.53-1.15; moderate heterogeneity between studies). However, an association was observed in the European descent groups (OR = 1.5; 95%CI = 1.15-1.96; P = 0.003; no significant heterogeneity between studies). There was no evidence of publication bias according to funnel plot and the Egger regression test (a = -2.16, P = 0.23). There was

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heterogeneity between studies and no clear evidence of an association between the TGF- β 1 gene -C509T polymorphism and susceptibility to IgA nephropathy in the worldwide population. Subgroup analysis suggests that the TGF- β 1 gene -C509T polymorphism would not be a risk factor for IgA nephropathy in Asians but might play a role in Europeans. More studies are required for definitive conclusions.

Key words: IgA nephropathy; -C509T polymorphism; TGF-β1 gene; Meta-analysis

INTRODUCTION

Immunoglobulin A nephropathy (IgAN), the most common glomerulonephritis throughout the world, exhibits an indolent but slowly progressive course leading to end-stage renal failure in 30 to 40% of patients older than 30 years (Brezzi et al., 2009). Its variable clinical picture likely depends on the diverse interplay of environmental and genetic factors in each patient. During the past decade, several studies have investigated the polymorphisms of the genes related to molecules in the renin-angiotensin system - interleukin (IL)-1, IL-6 and IL-10 (Bantis et al., 2004), IL-1 receptor, tumor necrosis factor alpha, uteroglobin, T-cell receptors, and nephrin (Narita et al., 2003; Kovacs et al., 2006) - and data have emerged to support the hypothesis that genetic factors affect susceptibility to both the onset and progression of IgAN (Narita et al., 2002; Syrjanen et al., 2002; Wiwanitkit, 2006; Yamamoto et al., 2012).

Members of the transforming growth factor- β (TGF- β) family play a crucial role in the pathogenesis of renal fibrosis thanks to their prosclerotic activity exerted by inducing cell proliferation and extracellular matrix molecule deposition. TGF- β and its respective receptors are abnormally expressed in experimental and human nephropathies (Julian et al., 2007; Bhowmik et al., 2011). TGF- β 1 is a TGF- β isoform that is highly expressed during the onset and progression of various renal diseases including IgAN (Song et al., 2003; Hohenstein et al., 2008; Iwano, 2010; Lee, 2011). Furthermore, a close relationship between increased renal tissue levels of TGF- β 1 and fibronectin, one of the major TGF- β 1-regulated extracellular components, has been reported in IgAN patients (Lee, 2011). The TGF- β 1 gene is located on chromosome 19q13 and is highly polymorphic. Five polymorphisms in white populations have been identified: 2 in the promoter region at positions -800 and -509, 1 at position +72 in a non-translated region, and 2 in the signal sequence at positions +869 and +915. Among these polymorphisms, -C509T is reportedly associated with some clinical phenotypes (Awad et al., 1998; Baan et al., 2000; Yokota et al., 2000; Yamada et al., 1998, 2001; Lacha et al., 2001). Grainger et al. (1999) have found that the C-509T polymorphism is associated with the circulating concentration of TGF- β 1, which is significantly lower in white women with the CC genotype compared with that in subjects with other genotypes.

The function and location of the TGF- β 1 -C509T polymorphism make it a solid candidate for association with IgAN, so a candidate gene analysis of TGF- β 1 -C509T was recently performed. Functional variants of the gene encoding TGF- β 1 -C509T were reportedly associated with IgAN. In addition, Vuong et al. (2009) have reported an association between the TGF- β 1 -C509T polymorphism and IgAN. Lim et al. (2005) and Qin et al. (2008) have found a similar association. Thus, the TGF- β 1 -C509T susceptibility gene has an important role in the pathogenesis of IgAN.

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Despite the association between TGF- β 1 -C509T and IgAN observed in some Swedish, South Korea, and Chinese populations (Lim et al., 2005; Vuong et al., 2009; Li et al., 2011), the results of later studies are inconsistent (Carturan et al., 2004; Sato et al., 2004; Xue et al., 2005; Qin et al., 2008; Brezzi et al., 2009). Considering the possible small effect size of this genetic polymorphism in IgAN and the relatively small sample size in each study, discrepancies are likely to become apparent because a single study may have been underpowered for the detection of a small but real association.

Given the amount of accumulated data now available, a quantitative synthesis of the evidence using rigorous methods is important. The aim of this study was to assess the association of the TGF- β 1 -C509T polymorphism with the risk of IgAN by conducting a metaanalysis of individual datasets from all eligible case-control studies published to date.

MATERIAL AND METHODS

Data

PubMed, the Chinese Biomedical Database (CBM), Chinese National Knowledge Infrastructure, and Wanfang (Chinese) were searched for genetic association studies published before January 2012 that evaluated the TGF- β 1 -C509T polymorphism and IgAN in humans. We also reviewed the reference lists of relevant articles and performed Baidu and Google searches to identify additional studies. The PubMed search was run using the mesh terms ["Glomerulonephritis, IgA" (Mesh), "IgA nephropathy", "IgAN"], and ("TGF-1" or "TGF" or "TGF- β 1 -C509T" or "transforming growing factor"). In the Chinese Biomedical Database, China National Knowledge Infrastructure, and Wanfang, the following words were used: ("TGF- β 1" or "Chinese technical term of TGF- β 1") and ("IgA nephropathy" or "glomerulonephritis" or "relevant Chinese technical terms").

Study selection

The following criteria were used to identify relevant published studies: 1) the studies used case-control or cohort designs that determined the distribution of the C677T and A1298C genotypes in cancer-free subjects and subjects with gastric cancer diagnosed with histopathological biopsy, 2) Ethics Committee approval and informed patient consent were obtained, 3) the studies provided relative risk estimates and respective variance or the relevant information needed to calculate it, and 4) the distribution of the genotypes in control groups was in Hardy-Weinberg equilibrium (HWE), and 5) raw data were available for retrieval. When multiple publications were found for the same population, only the latest or largest study was included. Meeting abstracts, case reports, editorials, and review articles were excluded.

Data extraction and synthesis

Data were independently extracted by 2 reviewers (Heng Wang and Peng Li) using a standardized data extraction form. Discrepancies were resolved through discussion, and if consensus was not reached, the decision was made by a third reviewer (Zhan-chun Feng). The title and abstract of all potentially relevant articles were screened to determine their relevance.

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Full articles were also scrutinized if the title and abstract were ambiguous. The following information was collected from each study: first author's surname, year of publication, ethnicity, criteria of enrolled patients, study design, total number of cases and controls, and number of cases and controls with the -C509T polymorphism (TT vs CT and CC genotypes).

Statistical analysis

Because case-control studies were involved, odds ratio (ORs) were used to assess the strength of the association between the TGF- β 1 -C509T polymorphism and IgAN. We calculated the ORs and respective 95% confidence intervals (95%CIs) by comparing the carriers of rare alleles with the wild homozygote as CT + CC vs TT.

The chi-square test was used to check for HWE of genotypes in the control group of each study. Statistical heterogeneity among studies was assessed with Q and I² statistics (Higgins and Thompson, 2002). I² values of 25, 50, and 75% were considered to be low, moderate, and high estimates, respectively. Heterogeneity was considered to be significant at a P value of <0.10. Lower heterogeneity indicated a more credible result.

A fixed-effects model using the Mantel-Haenszel method and a random-effects model using the DerSimonian and Laird method were used to pool the results (Petitti, 1994). Random effects are more appropriate when heterogeneity is present compared with the fixed-effects model that was used without heterogeneity. The significance of the pooled OR was determined using the Z-test and considered to be significant at a P value of <0.05.

Publication bias was investigated using a funnel plot, in which the standard error of log (OR) of each study was plotted against its OR. An asymmetric plot suggested possible publication bias. Funnel plot asymmetry was assessed with the Egger linear regression test method, which uses a linear regression approach to measure funnel plot asymmetry on the natural logarithm scale of the OR (Egger et al., 1997). The result was credible when publication bias was absent.

Analyses were performed using Stata version 10 (StataCorp LP, College Station, TX, USA), Review Manager 4.2 (Cochrane Collaboration, http://www.cc-ims.net/RevMan). All P values were 2 sided. A P value of <0.05 was considered to be statistically significant.

RESULTS

Subject characteristics

Eight association studies relating to the TGF- β 1 gene -C509T polymorphism with susceptibility to IgAN met the inclusion requirements for the meta-analysis (Carturan et al., 2004; Sato et al., 2004; Lim et al., 2005; Xue et al., 2005; Qin et al., 2008; Brezzi et al., 2009; Vuong et al., 2009; Li et al., 2011). A total of 1364 cases and 1483 controls were investigated. Three studies were from China (Xue et al., 2005; Qin et al., 2008; Li et al., 2011), 2 from Italy (Carturan et al., 2004; Brezzi et al., 2009), 1 from Japan (Sato et al., 2004), 1 from Korea (Lim et al., 2005), and 1 from Sweden (Vuong et al., 2009). All cases of IgAN were diagnosed using World Health Organization diagnosis criteria. Selected characteristics of the 8 case-control studies of the relationship between the TGF- β 1 gene-C509T polymorphism and susceptibility to IgAN are summarized in Table 1.

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Table 1. Characteristics of published studies about the association between the TGF- β 1 -C509T polymorphism with immunoglobulin A nephropathy included in the meta-analysis and results of Hardy-Weinberg test for genotype distribution of the TGF- β 1 -C509T polymorphism in control groups of reviewed studies.

Investigator	Year	Country	Ethnicity	TGF-β1 -509 C/T polymorphism				HWE P
				Case		Control		
				CC+CT	TT	CC+CT	TT	
Carturan et al.	2004	Italy	European	84	11	83	21	0.193
Li et al.	2011	China	Asian	27	11	32	10	0.768
Vuong et al.	2009	Sweden	European	121	88	223	245	0.711
Xue et al.	2005	China	Asian	319	68	164	38	0.879
Lu et al.	2008	China	Asian	79	40	93	23	0.776
Sato et al.	2004	Japan	Asian	263	66	233	64	0.31
Lim et al.	2005	Korea	Asian	49	33	42	13	0.34
Brezzi et al.	2009	Italy	European	82	23	146	54	0.481

HWE = Hardy-Weinberg equilibrium.

The genotype frequencies in the control group were consistent with HWE (Table 2). The distribution of ORs from individual studies with respect to their standard deviation was symmetrical in the funnel plot (Figure 1). The Egger test was performed to provide statistical evidence of funnel plot symmetry (a = -2.16, t = -1.31, P = 0.237; Figure 2). These data provided no significant evidence of publication bias.

Table 2. Egger linear regression test for publication bias of the IgAN gene TGF- β 1 -C509T polymorphism (a =
-2.16, t = -1.31 , P = 0.23 for C allele vs T allele).

Egger test (standard efficiency)	Coefficient	Standard error t p>		p> t	95% Confidence interval	
Slope Bias	0.6235989	0.422415	1.48	0.190	-4.100134 -6.177464	1.657211 1.859879
Dias	-2.130/93	1.042344	-1.31	0.237	-0.1//404	1.0390/9

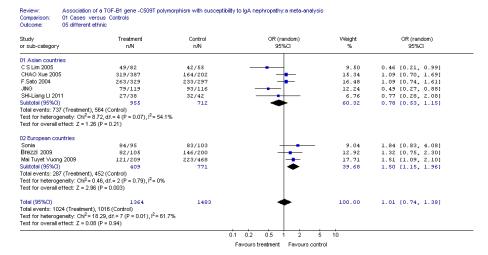


Figure 1. Funnel plot of 8 studies and relationship between TGF- β 1 -C509T and IgAN to determine publication bias. x-axis = log odds ratio (OR) of the TGF- β 1 -C509T allele; y-axis = standard error of log OR; d.f. = degrees of freedom; 95%CI = 95% confidence interval.

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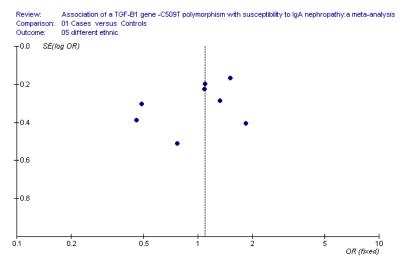


Figure 2. Egger linear regression test for publication bias of the IgAN gene TGF- β 1 -C509T polymorphism (a = -2.16, t = -1.31, P = 0.23 for C allele *vs* T allele). SE = standard error; OR = odds ratio.

Association between TGF-B1 -C509T and IgAN

Heterogeneity was observed among individual estimates of the ORs (chi-square = 18.29, P = 0.01), and the original data were combined by means of the random-effects model. Statistics calculated for each study are shown in a forest plot (Figure 3). The summary OR was 1.01 (95%CI = 0.74-1.38; P = 0.94) according to the random-effects model. No evidence was found to suggest that the TGF- β 1 gene -C509T polymorphism increased susceptibility to IgAN in a worldwide population.

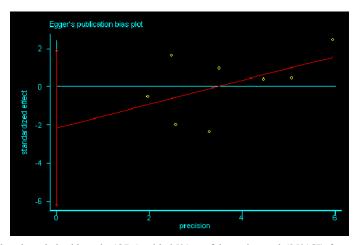


Figure 3. Overall and pooled odds ratio (ORs) with 95% confidence interval (95%CI) for overall analysis and subgroup analysis testing association of the TGF- β 1 -C509T polymorphism with IgAN. The aggregate OR and 95%CI of the risk allele are also given. The weighting factors (weight %) used to calculate the aggregate OR, calculated from the inverse of the variance, are given for each study.

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A previous report was published on the association between TGF- β 1 T869C and rheumatoid arthritis and revealed discrepancies that may have been due to ethic differences. What is meant by "discrepancies on the association?" What kinds of discrepancies? To look for an ethnic effect, we performed a subgroup meta-analysis in populations of Asian and European decent. The results indicated that the TGF- β 1 gene -C509T polymorphism had no effect on susceptibility to IgAN in subgroups of Asian descent (5 comparisons; OR = 0.78; 95%CI = 0.53-1.15; moderate between-study heterogeneity) under the random-effects model. However, an association was observed in the European subgroup under the fixed-effect models (OR = 1.5; 95%CI = 1.15-1.96; P = 0.003; no significant between-study heterogeneity; see Figure 3).

DISCUSSION

Since the first positive association reported between the TGF- β 1 gene -C509T polymorphism and susceptibility to IgAN in a Japanese population, 7 other studies have been undertaken to replicate the association. However, subsequent studies in other populations have been inconsistent. Therefore, we did a meta-analysis to estimate the relationship between the TGF- β 1 gene -C509T polymorphism and susceptibility to IgAN. Our meta-analysis is the first to examine this relationship. Overall, the results showed that the TGF- β 1 gene -C509T polymorphism may not be an IgAN susceptibility gene across populations. The common OR for the risk allele was 1.01 (95%CI = 0.74-1.38; P = 0.94). An association was found between the TGF- β 1 gene -C509T polymorphism and IgAN in European, but not Asian, populations in the subgroup meta-analysis.

Several explanations are possible for varying roles played by the same polymorphism in different ethnic populations or across different studies. First, the existence of the genetic heterogeneity results from the geographical environment and living habits of subject, which influence the frequency distribution of the crowd gene. Some studies have shown that the difference in C509T gene distribution between European and Asian countries is statistically significant (Cotton et al., 2002; Holla et al., 2002; Chen et al., 2005). Second, the clinical characteristics of participants - for example, age and years from onset, male predominance - may lead to different outcomes. Sometimes association can only be found in stratification analysis according to clinical characters. In this study, such information could not be obtained completely. Moreover, only 8 studies were included, which is inconvenient for subgroup studies. These influences may be the main source of the reported heterogeneity in this article.

The focus of meta-analysis is to combine comparable studies to increase sample size and statistical power and draw more compelling results. However, meta-analysis has confounding factors such as publication bias, diverse genetic backgrounds of subjects, and varying data-gathering protocols, methods of sampling, and quality analyses. A funnel plot and the Egger linear regression test used to assess publication bias suggested that bias was absent. Testing HWE for distribution of the genotypes in control groups also suggested that no significantly different genetic background was present among the participants. We followed the inclusion and exclusion criteria strictly according to standard protocols. The 8 studies appear to be comparable related to the meta-analysis.

Nevertheless, the analysis has several limitations. First, the number of total samples in the meta-analysis is small, and the amount of ethnic stratification in the subgroup analyses is even smaller. The amount of gender stratification is too small to conduct the subgroup analysis

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by gender. More studies are required to draw definite conclusions. Second, heterogeneity exists between studies, possibly owing to covariates: age and years from onset, gender, fraction of patients among these studies, and so on. Third, the selection of control participants in casecontrol studies may influence the results because hospital-based controls may not be as representative as population-based normal controls, although no evidence of an effect of the control population was detected in this meta-analysis. Considering the limited studies and population numbers included in our study, the results should be interpreted with caution.

CONCLUSIONS

Despite these limitations, this meta-analysis suggests that heterogeneity occurred between studies and found no clear evidence of an association between the TGF- β 1 gene -C509T polymorphism and susceptibility to IgAN in a worldwide population. Subgroup analysis results suggested that TGF- β 1 gene -C509T would not be a risk factor for IgAN in Asians but might play a role in IgAN susceptibility in Europeans. Because the number of studies included was relatively small, studies with larger pools of data are required for definitive conclusions.

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