

# Genetic parameter estimates and response to selection for weight and testicular traits in Nelore cattle 

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#### Abstract

We estimated genetic parameters for various phases of body and testicular growth until 550 days of age in Nelore cattle, using Bayesian inference, including correlation values and error estimates. Weight and scrotal records of 54,182 Nelore animals originating from 18 farms participating in the Brazilian Nelore Breeding Program (PMGRN) were included. The following traits were measured: weight at standard ages of 120 (W120), 210 (W210), 365 (W365), 450 (W450), and 550 (W550) days; weight gain between 120/210 (WG1), 210/365 (WG2), 365/450 (WG3), 450/550 (WG4), 120/365 (WG5), 120/450 (WG6), 120/550 (WG7), 210/450 (WG8), 210/550 (WG9), and 365/550 (WG10)


days of age; scrotal circumference at 365 (SC365), 450 (SC450) and 550 (SC550) days of age, and testicular growth between 365/450 (TG1), 450/550 (TG2) and 365/550 (TG3) days of age. The model included contemporary group (current farm, year and two-month period of birth, sex, and management group) and age of dam at calving, divided into classes as fixed effects. The model also included random effects for direct additive, maternal additive and maternal permanent environmental, and residual effects. The direct heritability estimates ranged from 0.23 to $0.39,0.13$ to 0.39 and 0.32 to 0.56 for weights at standard ages, weight gains and testicular measures, respectively. The genetic correlations between weights ( 0.69 to 0.94 ) and scrotal circumferences ( 0.91 to 0.97 ) measured at standard ages were higher than those between weight gain and testicular growth ( 0.18 to 0.97 and 0.36 to 0.77 , respectively). The weights at standard ages responded more effectively to selection, and also gave strong correlations with the other traits.

Key words: Beef cattle; Gibbs samples; Growth; Reproductive traits

## INTRODUCTION

The new economic order, based on competition and a constant search for productive efficiency, has forced livestock producers to adopt new technologies. These changes in farm management are accompanied by genetic alterations of the herds promoted by the selection of animals. This process is supported by the choice of criteria compatible with the objective of the activity and by the accurate estimation of (co)variance components of the traits selected.

In beef cattle, weights and weight gains between certain ages have been adopted as selection criteria because they are easy to measure and interpret, the possibility of promoting substantial genetic gain, and high genetic correlations between these measures. In Brazil, genetic breeding programs include scrotal circumference as a reproductive criterion since, in addition to its easy measurement, this trait presents favorable genetic correlations with productive and reproductive traits in both females (Eler et al., 2004; Forni and Albuquerque, 2005) and males (Sarreiro et al., 2002; Kealey et al., 2006).

The estimation of (co)variance components for any criteria adopted permits the construction and application of selection indices and of mixed model equations, as well as the calculation of parameters such as heritability and genetic correlation, thus facilitating the planning of breeding programs and the interpretation of genetic mechanisms involved in the expression of different traits (Henderson, 1986). In view of this importance, various methods of estimation have been tested in quantitative genetic analyses to obtain accurate estimates and to facilitate the implementation of complex models. In this respect, Gianola and Fernando (1986) introduced the Bayesian approach in animal breeding and Wang et al. (1993) subsequently applied it to linear mixed models using the Gibbs algorithm.

The objective of the present study was to test linear mixed models in combined analyses by Bayesian inference to estimate genetic parameters for different phases of body and testicular growth until 550 days of age in Nelore cattle, to contribute to selection programs designed to improve these traits.

## MATERIAL AND METHODS

Weight and scrotal records of Nelore animals originating from 18 farms participating in the Brazilian Nelore Breeding Program (PMGRN) were used. The following traits were studied: weight at standard ages of 120 (W120), 210 (W210), 365 (W365), 450 (W450), and 550 (W550) days; weight gain between 120/210 (WG1), 210/365 (WG2), 365/450 (WG3), 450/550 (WG4), 120/365 (WG5), 120/450 (WG6), 120/550 (WG7), 210/450 (WG8), 210/550 (WG9), and 365/550 (WG10) days of age; scrotal circumference at 365 (SC365), 450 (SC450) and 550 (SC550) days of age, and testicular growth between 365/450 (TG1), 450/550 (TG2) and 365/550 (TG3) days of age.

Exploratory analysis was performed using the Statistical Analysis System software (SAS ${ }^{\circ}$ Institute, 2003) to evaluate the consistency of the data. The contemporary groups consisted of current farm, year and two-month period of birth, sex, and management group. Records that were 3.5 standard deviations above or below the mean of the group and contemporary groups containing progeny from a single sire or fewer than five animals were eliminated. The file contained data from 54,182 individuals (Table 1) and the number of animals in the relationship matrix was 82,692 .

Table 1. Summary of data structure and descriptive statistics for weight at standard ages (W), weight gain (WG), scrotal circumference (SC), and testicular growth (TG) in Nelore cattle.

| Traits | N | Mean | SD | Min. | Max. | NS | NM |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| W120 (kg) | 50,956 | 127.60 | 19.07 | 57.00 | 199.00 | 1,090 | 21,576 |
| W210 (kg) | 47,772 | 185.64 | 27.88 | 78.00 | 312.00 | 1,097 | 20,723 |
| W365 (kg) | 39,588 | 238.83 | 40.14 | 101.00 | 452.00 | 1,064 | 18,439 |
| W450 (kg) | 36,540 | 278.84 | 47.66 | 118.00 | 559.00 | 1,023 | 17,402 |
| W550 (kg) | 23,146 | 319.82 | 53.69 | 142.00 | 606.00 | 808 | 13,158 |
| WG1 (kg) | 45,139 | 57.74 | 14.09 | -2.00 | 130.00 | 1,061 | 19,936 |
| WG2 (kg) | 35,237 | 52.37 | 25.27 | -41.00 | 183.00 | 988 | 16,925 |
| WG3 (kg) | 31,917 | 40.73 | 16.03 | -37.00 | 128.00 | 979 | 15,805 |
| WG4 (kg) | 20,738 | 46.88 | 20.15 | -23.00 | 133.00 | 765 | 12,084 |
| WG5 (kg) | 37,907 | 110.44 | 30.42 | -4.00 | 290.00 | 1,008 | 17,795 |
| WG6 (kg) | 33,980 | 151.97 | 37.07 | 16.00 | 402.00 | 954 | 16,359 |
| WG7 (kg) | 21,260 | 195.22 | 44.05 | 58.00 | 444.00 | 750 | 12,152 |
| WG8 (kg) | 33,679 | 93.71 | 33.75 | -46.00 | 279.00 | 953 | 16,326 |
| WG9 (kg) | 21,139 | 137.06 | 39.18 | 5.00 | 358.00 | 747 | 12,112 |
| WG10 (kg) | 21,774 | 85.43 | 28.44 | -64.00 | 261.00 | 794 | 12,444 |
| SC365 (cm) | 12,428 | 20.26 | 2.02 | 12.10 | 30.30 | 713 | 8,665 |
| SC450 (cm) | 14,715 | 23.43 | 2.81 | 12.70 | 36.30 | 736 | 9,640 |
| SC550 (cm) | 7,896 | 26.51 | 3.22 | 13.50 | 38.60 | 533 | 6,013 |
| TG1 (cm) | 10,241 | 3.20 | 1.30 | -1.50 | 8.90 | 652 | 7,403 |
| TG2 (cm) | 6,717 | 3.44 | 1.34 | -1.40 | 9.10 | 486 | 5,216 |
| TG3 (cm) | 5,624 | 6.63 | 2.08 | -1.90 | 17.50 | 485 | 4,528 |

$\mathrm{N}=$ number of records; $\mathrm{SD}=$ standard deviation; $\mathrm{NS}=$ number of sires; $\mathrm{NM}=$ number of mothers.
(Co)variance components were estimated by multi-trait analysis simultaneously including the following traits: W120, SC365, WG1, WG2, WG3, WG4, TG1, and TG2. The model included contemporary group and age of dam at calving divided into classes (for W120 and WG1) as fixed effects. The model also includes random effects for direct additive, maternal additive (for W120 and WG1) and maternal permanent environmental (for W120 and WG1), and residual. In view of contradictory results in the literature, the covariance between direct and maternal genetic effects was fixed at zero as done in other studies (Albuquerque and Meyer, 2001a).

The GIBBS2F90 program (Misztal, 2007) was used for analysis. A chain of 1,131,500 cycles was generated, with a burn-in period of 30,000 iterations and a sampling interval of 100. Genetic parameters for the other characteristics (W205, W365, W450, W550, WG5, WG6, WG7, WG8, WG9, WG10, SC450, SC550, and TG3) were estimated through the sum of variances in each vector

The matrix presentation of the complete model used for analysis is:

$$
y=X \beta+Z_{1} a+Z_{2} m+W c+e
$$

(Equation 1)
where $y, \beta, a, m, c$, and $e$ are vectors of observations, fixed effects, direct additive genetic effects, maternal additive genetic effects, maternal permanent environmental effects, and residual effects, respectively, and $X, Z_{1}, Z_{2}$, and $W$ are incidence matrices that relate $\beta, a, m$, and $c$ to the observations. Uniform and Gaussian priors were assumed for fixed and random effects, respectively:

$$
\begin{gathered}
\beta \propto \text { constante } \\
a \mid G_{a} \sim M V N\left[0,\left(A \otimes G_{a}\right)\right] \\
m \mid G_{m} \sim M V M\left[0,\left(A \otimes G_{m}\right)\right] \\
c \mid E_{p} \sim M V N\left[0,\left(I_{n} \otimes E_{p}\right)\right] \\
e \mid R \sim M V N\left[0,\left(I_{n} \otimes R\right)\right]
\end{gathered}
$$

(Equation 2)
where $A, G_{\mathrm{a}}, G_{\mathrm{m}}, E_{\mathrm{p}}, R$, and $I_{\mathrm{n}}$ are the relationship matrices, direct genetic, maternal genetic, maternal permanent environmental and residual covariances, and identity, respectively, and $\otimes$ is the operator of the Kronecker product. For variance components, priors derived from inverse Wishart distributions, corresponding to the inverse chi-square distribution in univariate cases, were used (Van Tassell and Van Vleck, 1996). Thus:

$$
\begin{gather*}
G_{a} \mid S_{a}, v_{a} \sim I W\left[S_{a} v_{a}, v_{a}\right] \\
G_{m} \mid S_{m}, v_{m} \sim I W\left[S_{m} v_{m}, v_{m}\right] \\
E_{p} \mid S_{p}, v_{p} \sim I W\left[S_{p} v_{p}, v_{p}\right]  \tag{Equation3}\\
R \mid S_{r}, v_{r} \sim I W\left[S_{r} v_{r}, v_{r}\right]
\end{gather*}
$$

where $S_{\mathrm{a}}$ and $v_{\mathrm{a}}, S_{\mathrm{m}}$ and $v_{\mathrm{m}}, S_{\mathrm{p}}$ and $v_{\mathrm{p}}$, and $S_{\mathrm{r}}$ and $v_{\mathrm{r}}$ are a priori values and degrees of freedom. for direct additive, maternal additive, maternal permanent environmental, and residual covariances, respectively.

The burn-in period and the sampling interval were established empirically and postGibbs analysis was performed using the Gibanal program (Van Kaam, 1997) to determine convergence and the number of effective samples. The population parameters were calculated in each vector, and the mean, median and mode of posterior distributions were estimated. In addition, the $95 \%$ high-density interval (HDI) was calculated since this parameter is more informative than confidence intervals in the case of asymmetric or multimodal distributions (Hyndman, 1996).

## RESULTS AND DISCUSSION

## Analysis of convergence, dependence between samples and reliability of the estimates

The results obtained with the Gibanal program showed that the burn-in period was sufficient to determine convergence of the chain. The effective sample size obtained for the parameters was relatively small (22 to 1614) when compared to the total number of vectors generated. In general, the correlation estimates presented the greatest dependence, with a smaller effective sample size being associated with correlations of maternal origin.

Therefore, the use of a model simultaneously including a large number of traits, although representing a more extensive source of information (Schaeffer, 1984), increases the dependence between samples due to the number of parameters involved. In the case of estimates related to maternal effects, the number of offspring per dam may have increased the dependence between these parameters, although other factors such as the number of dams with records also interfere with these estimates (Heydarpour et al., 2008). However, no minimum effective sample size on which to infer has been reported in the literature. The size obtained in the present study seemed to be sufficient to obtain measures of central tendency and variation.

With respect to the reliability of the estimates, heritabilities and variances presented a small HDI and a low standard deviation, findings indicating that the parameters were reliable. The correlation estimates presented low standard deviations, but exceptions were observed.

## Genetic parameter estimates and response to selection

The direct heritability estimates for weights at standard ages were moderate to high ( 0.23 to 0.39 ) and were higher at older ages (Table 2). The magnitude of variation in these estimates was similar to that reported by Albuquerque and El Faro (2008), but the point estimates were slightly higher. The growing trend in the heritability estimates, as also demonstrated by Albuquerque and Meyer (2001b) using random regression, is due to the fact that direct additive variance increases at a higher rate than residual variance with increasing age as a result of the combined action of a large number of genes on weights measured at the ages studied, whereas random environmental effects tend to cancel each other over time.

With respect to weight gains, the direct heritability estimates ranged from 0.13 to 0.39 and tended to be higher for wider intervals (WG5 to WG9). Similar estimates have been reported in the literature. Paneto et al. (2002) obtained estimates ranging from 0.16 to 0.32 for total gains between standard ages. Krejcová et al. (2007), studying daily weight gains over nine periods by multi-trait and random regression analysis, reported heritabilities ranging from 0.045 to 0.29 .

Considering the results presented, the selection based on WG6 or WG7 was found to be as effective as selection based on weights at standard ages, because a substantial part of the variation was due to the average effect of genes. However, it is important to consider the magnitude of direct additive variance, its relationship with other traits of interest and operational aspects, because it requires at least two measures in animal.

With respect to maternal effects, the heritability estimates were of low magnitude and were higher for weights at standard ages ( 0.11 and 0.10 for W120 and W210, respectively) when compared to weight gain ( 0.04 for WG1). Albuquerque and Meyer (2001b) reported higher maternal heritabilities

Table 2. Descriptive statistics of posterior distribution for variance components and genetic parameters for weight at standard age (W) and weight gains (WG).

| Traits | Parameters | Mean | Mode | Median | SD | ES |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| W120 | $\sigma^{2}$ | 50.75 | 50.82 | 50.73 | 3.10 | 754 | 44.61 | 56.64 |
|  | $\sigma^{\text {a }}$ | 121.53 | 122.01 | 121.50 | 1.94 | 942 | 117.80 | 125.80 |
|  | $\sigma^{2}{ }^{\text {e }}$ | 23.42 | 23.00 | 23.27 | 2.23 | 377 | 19.61 | 27.78 |
|  | $\sigma^{2 m}$ | 25.72 | 25.75 | 25.71 | 1.84 | 210 | 22.25 | 29.32 |
|  | $\mathrm{h}^{\text {pm }}$ | 0.23 | 0.23 | 0.23 | 0.01 | 754 | 0.20 | 0.25 |
|  | $\mathrm{h}^{2}{ }^{\text {d }}$ | 0.11 | 0.11 | 0.11 | 0.01 | 377 | 0.09 | 0.12 |
|  | $\mathrm{c}^{2{ }^{\text {m }}}$ | 0.12 | 0.12 | 0.12 | 0.01 | 283 | 0.10 | 0.13 |
| W210 | $\sigma^{2}$ | 102.28 | 102.98 | 102.39 | 6.38 | 707 | 90.65 | 114.76 |
|  | $\sigma^{a^{2}}$ | 226.56 | 227.04 | 226.65 | 4.08 | 943 | 218.92 | 234.48 |
|  | $\sigma^{2}$ | 41.64 | 40.73 | 41.73 | 4.02 | 404 | 34.39 | 49.66 |
|  | $\sigma_{\text {pm }}{ }^{\text {m }}$ | 54.64 | 53.41 | 54.59 | 3.48 | 453 | 48.11 | 61.41 |
|  | $\mathrm{h}^{\text {pr }}{ }^{\text {m }}$ | 0.24 | 0.24 | 0.24 | 0.01 | 566 | 0.21 | 0.27 |
|  | $\mathrm{h}^{2}{ }^{\text {d }}$ | 0.10 | 0.10 | 0.10 | 0.01 | 404 | 0.08 | 0.12 |
|  | $\mathrm{c}^{\text {c }}{ }^{\text {m }}$ | 0.13 | 0.13 | 0.13 | 0.01 | 453 | 0.11 | 0.14 |
| W365 | $\sigma^{2}$ | 173.00 | 175.61 | 173.01 | 9.27 | 941 | 154.40 | 190.80 |
|  | $\sigma^{2^{2}}$ | 322.38 | 321.17 | 322.28 | 6.44 | 1,414 | 310.74 | 335.27 |
|  | $\mathrm{h}^{\text {2 }}{ }_{\text {d }}$ | 0.35 | 0.34 | 0.35 | 0.02 | 941 | 0.32 | 0.38 |
| W450 | $\sigma^{2}$ | 234.88 | 236.54 | 234.76 | 12.11 | 1,614 | 211.82 | 258.45 |
|  | $\sigma^{2}{ }^{2}$ | 409.15 | 413.70 | 409.49 | 8.33 | 1,414 | 392.12 | 424.36 |
|  | $h^{2}{ }_{\text {d }}$ | 0.36 | 0.36 | 0.36 | 0.02 | +942 | 0.34 | 0.40 |
| W550 | $\sigma^{2}$ | 336.42 | 332.64 | 336.06 | 16.43 | 1,615 | 302.91 | 368.34 |
|  | $\sigma^{\text {a }}{ }_{\text {a }}$ | 533.29 | 534.18 | 533.46 | 11.94 | 1,413 | 509.13 | 556.31 |
|  | $\mathrm{h}^{\text {2 }}$ | 0.39 | 0.39 | 0.39 | 0.02 | 1,413 | 0.35 | 0.42 |
| WG1 | $\sigma^{2}$ | 18.21 | 18.20 | 18.20 | 1.37 | 566 | 15.66 | 20.70 |
|  | $\sigma^{\sigma^{2}}$ | 79.76 | 79.43 | 79.76 | 1.01 | 754 | 77.88 | 81.88 |
|  | $\sigma^{\text {c }}{ }^{\text {e }}$ | 4.47 | 4.58 | 4.46 | 0.67 | 147 | 3.31 | 5.77 |
|  | $\sigma^{2 m}$ | 5.88 | 5.89 | 5.87 | 0.68 | 101 | 4.51 | 7.16 |
|  | $\mathrm{h}^{\text {pm }}$ | 0.17 | 0.17 | 0.17 | 0.01 | 566 | 0.14 | 0.19 |
|  | $\mathrm{h}^{2}{ }_{\mathrm{m}}^{\text {d }}$ | 0.04 | 0.04 | 0.04 | 0.01 | 147 | 0.03 | 0.05 |
|  | $\mathrm{c}^{\text {m }}$ | 0.05 | 0.05 | 0.05 | 0.01 | 101 | 0.04 | 0.07 |
| WG2 | $\sigma_{\text {a }}{ }^{\text {a }}$ | 40.54 | 39.57 | 40.52 | 2.77 | 565 | 34.52 | 45.40 |
|  | $\sigma^{\text {a }}{ }_{\text {a }}$ | 170.58 | 170.81 | 170.55 | 2.41 | 471 | 165.70 | 175.20 |
|  | $h^{2}{ }_{\text {d }}$ | 0.19 | 0.19 | 0.19 | 0.01 | 565 | 0.17 | 0.22 |
| WG3 | $\sigma^{2}$ | 16.56 | 16.50 | 16.53 | 1.53 | 377 | 13.79 | 19.52 |
|  | $\sigma^{2}{ }^{\text {a }}$ | 115.51 | 115.29 | 115.50 | 1.44 | 707 | 112.90 | 118.50 |
|  | $\mathrm{h}^{2}{ }_{\text {d }}$ | 0.13 | 0.12 | 0.13 | 0.01 | 377 | 0.10 | 0.15 |
| WG4 | $\sigma^{2}$ |  |  |  |  | 323 |  | 30.30 |
|  | $\sigma^{2}{ }^{\text {a }}$ | 152.87 | 152.80 | 152.80 | 2.35 | 566 | 148.20 | 157.60 |
|  | $\mathrm{h}^{2}{ }_{\text {d }}$ | 0.15 | 0.14 | 0.15 | 0.01 | 323 | 0.12 | 0.17 |
| WG5 | $\sigma^{2}$ | 77.83 | 77.43 | 77.71 | 4.35 | 566 | 69.55 | 86.56 |
|  | $\sigma^{\sigma^{2}}$ | 215.81 | 216.97 | 215.81 | 3.40 | 942 | 208.90 | 222.58 |
|  | $\mathrm{h}^{2}{ }_{\text {d }}$ | 0.27 | 0.27 | 0.27 | 0.01 | 707 | 0.24 | 0.29 |
| WG6 | $\sigma^{2}$ | 126.39 | 126.53 | 126.47 | 6.69 | 943 | 113.99 | 139.83 |
|  | $\sigma^{2}{ }^{\text {a }}$ | 299.10 | 299.05 | 299.10 | 5.34 | 1,414 | 288.79 | 309.60 |
|  | $\mathrm{h}^{\text {2 }}{ }_{\text {d }}$ | 0.30 | 0.30 | 0.30 | 0.01 | 943 | 0.27 | 0.33 |
| WG7 | $\sigma^{2}$ | 205.48 | 204.93 | 205.36 | 10.53 | 1,131 |  | 227.37 |
|  |  | 403.41 | 402.53 | 403.27 | 8.56 | 1,414 | 386.85 | 420.05 |
|  | $\mathrm{h}^{\text {2 }}{ }_{\text {d }}$ | 0.34 | 0.33 | 0.34 | 0.02 | 1,131 | 0.31 | 0.37 |
| WG8 |  | 82.45 | 80.99 | 82.28 | 4.97 | 754 | 72.70 | 91.93 |
|  | $\sigma^{2}{ }^{\text {a }}$ | 263.94 | 262.55 | 263.88 | 4.17 | 1,131 | 255.40 | 271.67 |
|  | $h^{2}{ }_{\text {d }}$ | 0.24 | 0.24 | 0.24 | 0.01 | 942 | 0.21 | 0.26 |
| WG9 | $\sigma_{\text {a }}{ }^{1}$ | 139.63 | 139.13 | 139.47 | 7.97 | 942 | 123.76 | 155.16 |
|  | $\sigma^{2}{ }_{\text {a }}$ | 347.49 | 346.56 | 347.26 | 6.63 | 942 | 334.64 | 361.34 |
|  | $\mathrm{h}^{\text {c }}{ }_{\text {d }}$ | 0.29 | 0.29 | 0.29 | 0.02 | 1,131 | 0.26 | 0.32 |
| WG10 | $\sigma^{2}$ | 53.34 | 52.58 | 53.27 | 4.15 | 538 | 45.30 | 61.76 |
|  | $\sigma^{2{ }^{\text {a }}}$ | 247.14 | 248.05 | 247.07 | 3.94 | 754 | 240.06 | 255.67 |
|  | $\mathrm{h}^{\text {2 }}{ }_{\text {e }}$ | 0.18 | 0.18 | 0.18 | 0.01 | 377 | 0.15 | 0.20 |

$\sigma_{a}^{2}=$ direct additive variance; $\sigma_{\mathrm{e}}^{2}=$ residual variance; $\sigma_{\mathrm{m}}^{2}=$ maternal additive variance; $\sigma_{\mathrm{pm}}^{2}=$ maternal permanent environmental variance; $\mathrm{h}_{\mathrm{d}}{ }^{2}=$ direct heritability; $\mathrm{h}_{\mathrm{m}}^{2}=$ maternal heritability; $\mathrm{c}^{2}=$ maternal permanent environmental variance component as proportion of phenotypic variances; $\mathrm{ES}=$ number of effective samples; $\mathrm{SD}=$ standard deviation; $\mathrm{HDI}=$ high-density interval.
for weights at $120(0.25)$ and $240(0.20)$ days of age. In contrast to these authors who observed a decrease in the maternal effect with age, in the present study this effect was stable at 120 and 210 days. Therefore, in view of the similarity of maternal effects at these ages and the higher phenotypic variance observed at 210 days, a greater response to selection is expected for maternal ability at this age.

Regarding measures of testicular growth, the heritability estimates were high ( 0.32 to 0.56), except for TG2 (Table 3). As observed for the weight traits, greater response to selection can be achieved through selection based on the scrotal circumference at standard ages. For traits obtained at standard ages, the estimates were higher than those reported by Yokoo et al. (2007) ( $0.48,0.53$ and 0.42 for SC365, SC450 and SC550, respectively). Similar estimates for TG1 and TG2 were observed by Paneto et al. (2002) ( 0.24 and 0.18 for scrotal growth between $365 / 455$ and 455/550 days of age).

Table 3. Descriptive statistics of posterior distribution for variance components and genetic parameters for scrotal circumference (SC) and testicular growth (TG).

| Traits | Parameters | Mean | Mode | Median | SD | ES | HDI (95\%) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| SC365 | $\sigma^{2}$ | 1.52 | 1.52 | 1.52 | 0.09 | 629 | 1.34 | 1.70 |
|  | $\sigma^{2}{ }^{\text {a }}$ | 1.41 | 1.41 | 1.41 | 0.07 | 754 | 1.27 | 1.54 |
|  | $\mathrm{h}^{2}{ }_{\mathrm{d}}{ }^{\text {c }}$ | 0.52 | 0.51 | 0.52 | 0.03 | 629 | 0.47 | 0.57 |
| SC450 | $\sigma^{2}$ | 2.98 | 3.01 | 2.99 | 0.19 | 629 | 2.61 | 3.34 |
|  | $\sigma^{2}{ }_{\text {a }}$ | 2.36 | 2.29 | 2.36 | 0.14 | 628 | 2.12 | 2.63 |
|  | $h^{2}{ }_{\text {d }}$ | 0.56 | 0.56 | 0.56 | 0.03 | 706 | 0.50 | 0.61 |
| SC550 | $\sigma^{2}$ | 3.82 | 3.72 | 3.81 | 0.26 | 470 | 3.33 | 4.32 |
|  | $\sigma^{2}{ }_{\text {a }}$ | 3.06 | 3.10 | 3.07 | 0.19 | 377 | 2.67 | 3.43 |
|  | $h^{2}{ }_{\text {d }}$ | 0.56 | 0.58 | 0.56 | 0.03 | 471 | 0.49 | 0.62 |
| TG1 | $\sigma_{a}^{2}$ | 0.44 | 0.45 | 0.44 | 0.04 | 313 | 0.36 | 0.53 |
|  | $\sigma^{2}{ }_{\text {a }}$ | 0.90 | 0.91 | 0.90 | 0.04 | 376 | 0.84 | 0.98 |
|  | $h^{2}{ }_{\text {d }}$ | 0.33 | 0.33 | 0.33 | 0.03 | 313 | 0.27 | 0.38 |
| TG2 | $\overline{\sigma_{a}^{2}}$ |  |  | 0.25 | 0.05 | 140 | 0.18 | 0.36 |
|  | $\sigma^{2}{ }^{2}$ | 1.14 | 1.15 | 1.14 | 0.05 | 147 | 1.06 | 1.23 |
|  | $\mathrm{h}_{\mathrm{d}}^{\mathrm{c}^{\mathrm{e}}}$ | 0.18 | 0.18 | 0.18 | 0.03 | 113 | 0.12 | 0.25 |
| TG3 |  | 0.95 | 0.96 | 0.95 | 0.12 | 202 | 0.72 | 1.16 |
|  | $\sigma^{2}{ }_{\text {a }}{ }^{\text {a }}$ | 2.06 | 2.10 | 2.07 | 0.10 | 180 | 1.87 | 2.24 |
|  | $h^{2}{ }_{\text {d }}$ | 0.32 | 0.32 | 0.32 | 0.04 | 202 | 0.24 | 0.38 |

$\overline{\sigma_{\mathrm{a}}^{2}}=$ direct additive variance; $\sigma_{\mathrm{e}}^{2}=$ residual variance; $\mathrm{h}_{\mathrm{d}}^{2}=$ direct heritability; $\mathrm{SD}=$ standard deviation; $\mathrm{ES}=$ number of effective samples; $\mathrm{HDI}=$ high-density interval.

The genetic correlations between weights ( 0.69 to 0.94 ) and scrotal circumferences ( 0.91 to 0.97 ) measured at standard ages were higher than those between weight gain and testicular growth ( 0.18 to 0.97 and 0.36 to 0.77 , respectively) (Table 4 ). High correlations for weight gain and testicular growth were only observed between overlapping measures. Similar results have been reported by Albuquerque and Meyer (2001b) for weights at standard ages using random regression models and by Yokoo et al. (2007) between measures of scrotal circumference at 365,450 and 550 days of age. The genetic correlations obtained for weight gains showed a magnitude similar to that reported by Krejcová et al. (2007), except for some negative correlations ( -0.131 to 0.958 ).

For traits measured at standard ages (weight and scrotal circumference), an inverse relationship was observed between the genetic correlation estimate and chronological distance between the two measures. Similar results have been reported by Albuquerque and Meyer (2001b) using random regression models for weights. Weight gains presented higher correlations with weights adjusted for age at the upper limit of the interval that determined the weight gain.

|  | W120 | W210 | W365 | W450 | W550 | WG1 | WG2 | WG3 | WG4 | WG5 | WG6 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| W120 | - | 0.94 (0.01) | 0.78 (0.02) | 0.73 (0.02) | 0.69 (0.02) | 0.55 (0.04) | 0.12 (0.05) | 0.24 (0.05) | 0.29 (0.06) | 0.35 (0.04) | 0.36 (0.04) |
| W210 | 0.81 (0.00) | - | 0.88 (0.01) | 0.82 (0.01) | 0.80 (0.02) | 0.81 (0.02) | 0.24 (0.05) | 0.25 (0.05) | 0.42 (0.05) | 0.56 (0.03) | 0.53 (0.03) |
| W365 | 0.58 (0.01) | 0.70 (0.01) | - | 0.97 (0.00) | 0.94 (0.01) | 0.80 (0.02) | 0.66 (0.03) | 0.43 (0.05) | 0.47 (0.05) | 0.86 (0.01) | 0.83 (0.02) |
| W450 | 0.52 (0.01) | 0.61 (0.01) | 0.85 (0.00) |  | 0.97 (0.00) | 0.73 (0.03) | 0.70 (0.02) | 0.63 (0.04) | 0.48 (0.06) | 0.86 (0.01) | 0.90 (0.01) |
| W550 | 0.49 (0.01) | 0.59 (0.01) | 0.73 (0.01) | 0.85 (0.00) |  | 0.76 (0.03) | 0.67 (0.03) | 0.60 (0.04) | 0.68 (0.04) | 0.85 (0.01) | 0.88 (0.01) |
| WG1 | 0.13 (0.01) | 0.69 (0.01) | 0.47 (0.01) | 0.39 (0.01) | 0.39 (0.01) | - | 0.36 (0.05) | 0.18 (0.06) | 0.52 (0.06) | 0.74 (0.02) | 0.65 (0.03) |
| WG2 | -0.14 (0.01) | -0.19 (0.01) | -0.06 (0.01) | -0.05 (0.01) | 0.33 (0.01) | -0.15 (0.01) |  | 0.50 (0.05) | 0.31 (0.06) | 0.89 (0.01) | 0.88 (0.01) |
| WG3 | 0.01 (0.00) | -0.02 (0.01) | -0.07 (0.01) | 0.47 (0.01) | 0.37 (0.01) | -0.05 (0.01) | -0.08 (0.01) |  | 0.27 (0.07) | 0.44 (0.05) | 0.71 (0.03) |
| WG4 | 0.07 (0.01) | 0.11 (0.01) | -0.02 (0.01) | -0.06 (0.01) | 0.48 (0.01) | 0.09 (0.01) | -0.15 (0.01) | -0.08 (0.01) |  | 0.47 (0.05) | 0.47 (0.05) |
| WG5 | -0.05 (0.01) | 0.25 (0.01) | 0.79 (0.01) | 0.65 (0.01) | 0.53 (0.01) | 0.48 (0.01) | 0.80 (0.00) | -0.10 (0.01) | -0.08 (0.01) | - | 0.95 (0.01) |
| WG6 | -0.03 (0.01) | 0.20 (0.01) | 0.62 (0.01) | 0.84 (0.00) | 0.67 (0.01) | 0.37 (0.01) | 0.63 (0.01) | 0.54 (0.01) | -0.11 (0.01) | 0.79 (0.00) | - |
| WG7 | 0.02 (0.01) | 0.24 (0.01) | 0.53 (0.01) | 0.69 (0.01) | 0.88 (0.01) | 0.38 (0.01) | 0.45 (0.01) | 0.41 (0.01) | 0.52 (0.01) | 0.63 (0.01) | 0.79 (0.01) |
| WG8 | -0.10 (0.01) | -0.17 (0.01) | 0.41 (0.01) | 0.68 (0.01) | 0.50 (0.01) | -0.15 (0.01) | 0.75 (0.00) | 0.60 (0.01) | -0.17 (0.01) | 0.58 (0.01) | 0.86 (0.00) |
| WG9 | -0.04 (0.01) | -0.07 (0.01) | 0.34 (0.01) | 0.55 (0.01) | 0.76 (0.01) | -0.07 (0.01) | 0.56 (0.01) | 0.47 (0.01) | 0.51 (0.01) | 0.45 (0.01) | 0.67 (0.01) |
| WG10 | 0.07 (0.01) | 0.07 (0.01) | -0.06 (0.01) | 0.27 (0.01) | 0.63 (0.01) | 0.04 (0.01) | -0.17 (0.01) | 0.62 (0.01) | 0.73 (0.01) | -0.13 (0.01) | 0.28 (0.01) |
| SC365 | 0.32 (0.02) | 0.39 (0.02) | 0.52 (0.02) | 0.49 (0.02) | 0.44 (0.02) | 0.27 (0.02) | 0.26 (0.02) | 0.06 (0.02) | 0.02 (0.02) | 0.40 (0.02) | 0.37 (0.02) |
| SC450 | 0.30 (0.02) | 0.36 (0.02) | 0.48 (0.02) | 0.53 (0.02) | 0.52 (0.02) | 0.25 (0.02) | 0.24 (0.02) | 0.20 (0.02) | 0.09 (0.03) | 0.36 (0.02) | 0.44 (0.02) |
| SC550 | 0.31 (0.02) | 0.38 (0.03) | 0.46 (0.03) | 0.50 (0.02) | 0.57 (0.03) | 0.26 (0.02) | 0.20 (0.02) | 0.18 (0.02) | 0.25 (0.03) | 0.33 (0.02) | 0.39 (0.03) |
| TG1 | 0.08 (0.02) | 0.09 (0.02) | 0.13 (0.02) | 0.25 (0.02) | 0.28 (0.02) | 0.06 (0.02) | 0.07 (0.02) | 0.26 (0.02) | 0.12 (0.02) | 0.09 (0.02) | 0.24 (0.02) |
| TG2 | 0.08 (0.02) | 0.10 (0.02) | 0.07 (0.02) | 0.06 (0.02) | 0.20 (0.02) | 0.06 (0.02) | -0.02 (0.02) | -0.01 (0.02) | 0.27 (0.02) | 0.02 (0.02) | 0.02 (0.02) |
| T3 | 0.11 (0.02) | 0.13 (0.02) | 0.14 (0.02) | 0.21 (0.02) | 0.33 (0.02) | 0.09 (0.02) | 0.03 (0.02) | 0.17 (0.02) | 0.28 (0.02) | 0.08 (0.02) | 0.17 (0.02) |


|  | WG7 | WG8 | WG9 | WG10 | SC365 | SC450 | SC550 | TG1 | TG2 | TG3 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| W120 | 0.39 (0.04) | 0.19 (0.04) | 0.27 (0.04) | 0.34 (0.05) | 0.49 (0.04) | 0.41 (0.04) | 0.40 (0.04) | 0.16 (0.06) | 0.13 (0.09) | 0.18 (0.06) |
| W210 | 0.56 (0.03) | 0.27 (0.04) | 0.39 (0.04) | 0.43 (0.04) | 0.56 (0.03) | 0.49 (0.04) | 0.48 (0.04) | 0.24 (0.06) | 0.17 (0.09) | 0.25 (0.06) |
| W365 | 0.82 (0.02) | 0.66 (0.03) | 0.71 (0.03) | 0.57 (0.04) | 0.53 (0.03) | 0.49 (0.03) | 0.49 (0.04) | 0.28 (0.05) | 0.24 (0.08) | 0.32 (0.06) |
| W450 | 0.88 (0.01) | 0.77 (0.02) | 0.80 (0.02) | 0.68 (0.04) | 0.44 (0.03) | 0.41 (0.03) | 0.42 (0.04) | 0.25 (0.05) | 0.22 (0.09) | 0.28 (0.06) |
| W550 | 0.93 (0.01) | 0.74 (0.02) | 0.86 (0.01) | 0.81 (0.02) | 0.42 (0.04) | 0.38 (0.04) | 0.41 (0.04) | 0.22 (0.05) | 0.26 (0.08) | 0.29 (0.06) |
| WG1 | 0.69 (0.03) | 0.33 (0.05) | 0.48 (0.05) | 0.46 (0.05) | 0.50 (0.04) | 0.48 (0.04) | 0.47 (0.05) | 0.31 (0.06) | 0.16 (0.10) | 0.30 (0.07) |
| WG2 | 0.80 (0.02) | 0.92 (0.01) | 0.84 (0.02) | 0.49 (0.05) | 0.22 (0.05) | 0.23 (0.04) | 0.26 (0.05) | 0.19 (0.06) | 0.24 (0.09) | 0.26 (0.07) |
| WG3 | 0.65 (0.04) | 0.79 (0.03) | 0.73 (0.03) | 0.74 (0.03) | -0.07 (0.05) | -0.03 (0.05) | -0.02 (0.06) | 0.04 (0.07) | 0.02 (0.11) | 0.02 (0.09) |
| WG4 | 0.72 (0.03) | 0.34 (0.06) | 0.69 (0.03) | 0.84 (0.02) | 0.18 (0.06) | 0.15 (0.06) | 0.21 (0.06) | 0.05 (0.07) | 0.32 (0.10) | 0.20 (0.07) |
| WG5 | 0.91 (0.01) | 0.82 (0.02) | 0.84 (0.02) | 0.58 (0.05) | 0.40 (0.04) | 0.39 (0.04) | 0.41 (0.04) | 0.29 (0.05) | 0.25 (0.08) | 0.33 (0.06) |
| WG6 | 0.95 (0.01) | 0.93 (0.01) | 0.92 (0.01) | 0.72 (0.03) | 0.29 (0.04) | 0.30 (0.04) | 0.32 (0.04) | 0.24 (0.05) | 0.20 (0.09) | 0.27 (0.06) |
| WG7 | - | 0.85 (0.01) | 0.97 (0.00) | 0.86 (0.02) | 0.29 (0.04) | 0.29 (0.04) | 0.32 (0.04) | 0.20 (0.05) | 0.27 (0.08) | 0.28 (0.06) |
| WG8 | 0.64 (0.01) | - | 0.91 (0.01) | 0.68 (0.03) | 0.12 (0.04) | 0.14 (0.04) | 0.17 (0.05) | 0.16 (0.06) | 0.17 (0.08) | 0.20 (0.06) |
| WG9 | 0.90 (0.00) | 0.76 (0.01) | - | 0.88 (0.01) | 0.17 (0.04) | 0.18 (0.04) | 0.22 (0.04) | 0.14 (0.05) | 0.27 (0.08) | 0.24 (0.06) |
| WG10 | 0.69 (0.01) | 0.27 (0.01) | 0.72 (0.01) | - | 0.09 (0.05) | 0.09 (0.05) | 0.13 (0.05) | 0.06 (0.06) | 0.22 (0.09) | 0.16 (0.07) |
| SC365 | 0.33 (0.02) | 0.25 (0.02) | 0.23 (0.02) | 0.06 (0.02) | - | 0.96 (0.01) | 0.91 (0.01) | 0.63 (0.05) | 0.26 (0.08) | 0.57 (0.05) |
| SC450 | 0.43 (0.03) | 0.33 (0.02) | 0.34 (0.03) | 0.21 (0.02) | 0.78 (0.01) | - | 0.97 (0.01) | 0.83 (0.02) | 0.32 (0.08) | 0.74 (0.04) |
| SC550 | 0.49 (0.03) | 0.28 (0.02) | 0.41 (0.03) | 0.31 (0.02) | 0.58 (0.02) | 0.80 (0.01) | - | 0.83 (0.03) | 0.54 (0.07) | 0.85 (0.02) |
| TG1 | 0.28 (0.02) | 0.22 (0.02) | 0.27 (0.02) | 0.27 (0.02) | 0.02 (0.03) | 0.64 (0.02) | 0.56 (0.03) | - | 0.36 (0.09) | 0.87 (0.03) |
| TG2 | 0.18 (0.02) | -0.02 (0.02) | 0.17 (0.02) | 0.21 (0.02) | -0.19 (0.04) | -0.14 (0.04) | 0.49 (0.03) | 0.01 (0.03) | - | 0.77 (0.05) |
| TG3 | 0.32 (0.02) | 0.14 (0.02) | 0.30 (0.02) | 0.33 (0.02) | -0.12 (0.04) | 0.32 (0.03) | 0.74 (0.02) | 0.67 (0.02) | 0.75 (0.01) | - |

[^0]The highest residual correlations (Table 4) were observed between traits measured at standard ages, whereas the magnitude of residual correlations between measures of weight gain and testicular growth was low and negative correlations were frequently observed between subsequent measurements. Negative and low residual correlations were also observed between weight gain and weights and between testicular growth and measures of scrotal circumference. This finding indicates that environments favoring some traits can be unfavorable for others. These negative and low residual correlations especially between chronologically close measures suggest that climatic changes that occur during the year are the main factors responsible, although the contemporary groups included this effect.

With respect to weight and scrotal circumference, the highest residual correlation was observed between adjacent measures, i.e., the more distant the two measurements, the more distinct the environmental conditions and the non-additive genetic effects that influenced them. Analyzing the correlations between weight gains and weights at standard ages, the highest correlation was found between weight gain and weight measured at its upper limit, as observed for genetic correlations.

The higher heritabilities for weight and scrotal circumference measured at standard ages when compared to weight gain and scrotal growth also seem to be the result of the lower residual correlations compared to genetic correlations between gains. This fact also explains the increase of heritabilities for weight and scrotal circumference with increasing age of the animals.

Since weight and scrotal circumference measured at standard ages are normally used in Brazil as the main selection criteria for Nelore cattle, the response to selection of these traits was evaluated using different criteria that could be adopted (Figures 1 and 2). The use of weights at standard ages as selection criteria showed that, except for W210, direct selection is the most effective approach to increase weight at a certain age. Selection for weight at a certain age shows a similar efficiency to increase weights at previous ages as direct selection for these traits. WG1, WG6 and WG7 were found to be the most effective selection criteria among the weight gain traits. Among the testicular traits, SC365 was the most effective selection criterion to increase weights in the next generation.

When the objective was to increase scrotal circumference at different ages, the best selection criteria were W365 and WG5 among weight traits and TG1 among testicular traits since these traits showed the highest increases irrespective of age. With respect to scrotal circumference at a standard age, SC450 was found to be an attractive criterion since the gains obtained by indirect selection at other ages were similar to those obtained by direct selection.

## Considerations regarding the choice of selection criteria for Nelore cattle

First, weights at standard ages are the best criteria since they respond more effectively to selection, providing also substantial indirect gains in other traits. W550 was found to be the most effective selection criterion among the weight traits studied. However, the choice of the selection criterion should take into account the importance of the traits for the breeder and the cost of maintenance of the animals until the age of recording.

The estimated genetic correlations between weight gain and weight at a standard age suggest the alternative use of the former as a selection criterion since they provoke a smaller increase in birth weight and mature weight. In addition, other studies have demonstrated favorable associations between these measures and other traits of economic interest (Arthur et al., 2001). Further studies investigating weight gain as a selection criterion for Nelore cattle are necessary.


Figure 1. Expected response to selection for weights at standard ages, when the selection is done through: A. weights at standard ages (W, in kg ), B. weight gain (WG, in kg ), and C. scrotal circumference (SC) and testicular growth (TG), both in cm .


Figure 2. Expected response to selection for scrotal circumference, when the selection is done through: A. weights at standard ages (W, in kg ), B. weight gain (WG, in kg ), and C. scrotal circumference ( SC ) and testicular growth (TG), both in cm .

With respect to testicular traits, SC365 was found to be the best selection criterion since its heritability was similar to that of SC450 and SC550, in addition to presenting a higher genetic correlation with weights at standard ages. In this respect, the fact that this trait manifests at a younger age permits a reduction in the generation interval. According to Toelle and Robison (1985), early testicular traits are a reflex of body development, whereas late traits are influenced by the hormonal system. This statement is supported by the higher genetic correlations between weights and SC365 observed in this study, with higher testicular activity being observed after this age in Zebu animals. As a consequence, measures obtained after 365 days of age may present a higher correlation with female reproductive traits.

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[^0]:    $\mathrm{W}=$ weight at standard age; $\mathrm{WG}=$ weight gain

